# ENTREPRENEURSHIP CAPITAL AND REGIONAL PRODUCTIVITY REVISITED<sup>\*</sup>

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Government intervention to foster entrepreneurship has been justified by the positive spillovers that regional entrepreneurship capital could have on firms' productivity. The available evidence shows a positive relationship between entrepreneurship capital measures and regional production. This paper argues that this evidence could be explained not only by said spillovers but also by the presence of decreasing returns to scale in firms' production technology. The paper provides a simple methodological benchmark for distinguishing between and measuring the two effects. The analysis conducted using a sample of 52 Spanish provinces over 11 years confirms the presence of decreasing returns to scale. Therefore, the previous literature could have overestimated the spillovers of regional entrepreneurship capital on firms' production.

*Key words:* entrepreneurship capital, regional productivity, scale economies. *JEL classification:* R11, L26, O4.

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and significant effects of regional entrepreneurship capital on regional production, which are interpreted as evidence of positive spillovers or externalities.

Regional aggregated production functions have been extensively used by economists, but they are not free of limitations [see Fisher (1969, 2005) for further details]. One such limitation is the need to assume the existence of constant returns to scale. This is important in this case, because the most commonly used indicators of entrepreneurship capital have been based on the number of firms (incumbent or new, in absolute or relative terms, or their respective growth rate over time), so a positive correlation is expected between those entrepreneurship capital measures and the stock of firms in the region. The paper's contributions are based on the notion that a region's production is the aggregate of all production activities of the firms in the region. When the firms' technology has decreasing returns to scale (this is a necessary condition for the existence of a unique optimal firm size), the number of firms in the region will be positively related with the region's production level. In such cases, the spillovers of entrepreneurship capital on firms' productivity will be overestimated by the usual estimation procedures in the literature. In other words, given two regions that use the same levels of other inputs, a region with a higher number of firms could be more productive for two reasons: first, the technologies present decreasing returns to scale; or second, the entrepreneurship capital has positive spillovers or externalities on the firms in the region.

The main contribution of the paper is to provide a simple methodological benchmark to help distinguish between these two effects. The proposal is therefore to estimate the regional differences in firms' average production depending on the average use of inputs at the firm level (capital and labour) and the inputs available at the regional level (entrepreneurship capital and knowledge). This will provide estimates of the spillovers of entrepreneurship capital on the total factor productivity of the firms in the region. The analysis can be carried out with the standard data available in the literature. The only special requirement is information about the regional stock of firms.

The paper is organized as follows. First, we develop a theoretical framework to interpret the previous empirical literature and to discuss the methodological contributions proposed in this paper. Second, we apply that framework to a data sample covering Spain's 52 provinces (NUTS 3 Eurostat) in the 2002-2012 period with information as similar as possible to that used in the previously cited literature. Although previous studies have analysed the economic impact of entrepreneurship capital in Spain [Salas-Fumás and Sánchez-Asín (2008, 2010a, 2010b, 2013); Callejón (2009); Callejón and Ortún (2009)], those studies use other methodological approaches and aggregate data at the level of the Autonomous Communities (NUTS 2 Eurostat) into which provinces are grouped. As the effects of entrepreneurship capital seem to be stronger at the local level, smaller regional divisions are preferred when data is available. Finally, we discuss the paper's implications.

## 1. METHODOLOGICAL DISCUSSION

The concept of entrepreneurship capital and how to measure it have generated some discussion [Erikson (2002); Audretsch (2009); Bönte *et al.* (2008)], as is the case with any other kind of input. For example, empirical applications apply different

measures that range from the stock of firms in the region [Stough *et al.* (2008)] to the entry rate of firms in key industries [Chang (2011)]. Audretsch and Keilbach (2004a,b, 2008) used the annual average number of new firms with 1,000 workers created in a three-year period. Mueller (2006, 2007) also uses this indicator, along with the number of new firms created in one year. Sutter and Stough (2009) use the average number of technological and innovative firms created in the last five years, while Bönte *et al.* (2008), Salas-Fumás and Sánchez-Asín (2008, 2010a, 2010b, 2013) and Stough *et al.* (2008) use the regional self-employment rate. Inevitably, all of those entrepreneurship capital measures are (or could be) related with the number of firms in the region.

To estimate the impact of regional entrepreneurship capital on the production for region i and period t,  $Y_{i,t}$ , most of the literature follows Solow (1956). The inputs considered are capital ( $K_{i,t}$ ), labour ( $L_{i,t}$ ), knowledge ( $R_{i,t}$ ) and entrepreneurship capital ( $E_{i,t}$ ). The output obtained as a combination of those inputs is estimated in most cases by Cobb-Douglas (1928) functions:

$$\ln Y_{i,t} = \mu \ln R_{i,t} + \delta \ln E_{i,t} + \alpha \ln K_{i,t} + \beta \ln L_{i,t} + \varepsilon_{i,t}$$
<sup>[1]</sup>

Hence, the parameters to be estimated are production elasticity with respect to capital ( $\alpha$ ), labour ( $\beta$ ), entrepreneurship capital ( $\delta$ ) and knowledge ( $\mu$ ). As usual, the error terms ( $\varepsilon_{i,t}$ ) are independent and identically distributed and follow a standard normal distribution. Studies with panel data can control for the regional and time fixed effects. When it is assumed that production technologies present constant returns to scale for private inputs ( $\beta = 1 - \alpha$ ) the production per employee ( $y_{i,t} = Y_{i,t} / L_{i,t}$ ) will be:

$$\ln y_{i,t} = \ln Y_{i,t} - \ln L_{i,t} = \mu \ln R_{i,t} + \delta \ln E_{i,t} + \alpha \ln k_{i,t} + \varepsilon_{i,t}$$
[2]

where  $k_{i,t} = K_{i,t} / L_{i,t}$ . While Audretsch and Keilbach (2004a) take a panel data approach and estimate Equation [2], other studies directly have estimated Equation [1] [Audretsch and Keilbach (2004b); Audretsch *et al.* (2008); Mueller (2006, 2007); Bönte *et al.* (2008); Stough *et al.* (2008)]. In all studies that estimate Equation [1], with the exception of Audretsch and Keilbach (2004b), there are decreasing returns to scale ( $\beta + \alpha < 1$ ). Only Mueller (2006), however, reports a test of their significance, showing the elasticity of production with respect to knowledge ( $\mu$ ) and entrepreneurship capital ( $\delta$ ) to be positive and statistically significant.

In previous literature, the elasticity of production with respect to entrepreneurship capital ( $\delta$ ) has been interpreted as a measure of its spillovers or externalities. This requires some simplifying assumptions regarding the aggregation process of the firms' production functions at the regional level.

To show this, we consider  $Y_{i,t}$  to be the aggregation of the production of the  $n_{i,t}$  firms operating in region i during period t. Define  $Y_{i,t,j}$  as the production of firm j in region i during period t. In algebraic terms:

$$Y_{i,t} = \sum_{j=1}^{n_{i,j}} Y_{i,t,j} = n_{i,t} \overline{Y}_{i,t} \quad \text{where} \quad \overline{Y}_{i,t} = \sum_{j=1}^{n_{i,j}} Y_{i,t,j} / n_{i,t}$$
[3]

Firm j can use a set of private inputs purchased on the market. To reduce notation, and for the sake of consistent notation with previous literature, we only consider capital  $K_{i,t,j}$  and labour  $L_{i,t,j}$  as the inputs purchased by the firms, while entrepreneurship capital  $E_{i,t}$  and knowledge  $R_{i,t}$  are available to all the firms and inhabitants of region i. The following Cobb-Douglas (1928) function summarizes the relationship between the production and inputs used:

$$Y_{i,t,j} = R_{i,t}^{\mu} E_{i,t}^{\delta} K_{i,t,j}^{\alpha} L_{i,t,j}^{\beta}$$
[4]

The estimation of regional production by equations [3] and [4] requires information about the level of output and private inputs used by all the firms in the region. With aggregate data at the regional level, the best approximation to each firm output and private inputs is the firm average for this region,  $\overline{Y}_{i,j}$ ,  $\overline{K}_{i,j} = \sum_{j=1}^{n_{i,j}} K_{i,t,j}/n_{i,t}$ and  $\overline{L}_{i,j} = \sum_{j=1}^{n_{i,j}} L_{i,t,j}/n_{i,t}$ . So we can rewrite equation [4] as follows:

$$\overline{Y}_{i,t} = R^{\mu}_{i,t} E^{\delta}_{i,t} \overline{K}^{\alpha}_{i,t,j} \overline{L}^{\beta}_{i,t}$$
<sup>[5]</sup>

Given that the Cobb-Douglas is a homogenous function of degree  $\theta = \beta + \alpha$ ,  $f(\overline{K}, \overline{L}) = (\frac{K_{i,t}}{n_{i,t}})^{\alpha} (\frac{L_{i,t}}{n_{i,t}}) \beta = n_{i,t}^{-\theta} K_{i,t}^{\alpha} L_{i,t}^{\beta}$ , then from equations [3] and [5] we can obtain:

$$Y_{i,t} = n_{i,t}^{1-\theta} R_{i,t}^{\mu} E_{i,t}^{\delta} K_{i,t}^{\alpha} L_{i,t}^{\beta}$$
[6]

This approach can easily be implemented, but we have not found any estimations of Equation [5] or [6] in the reviewed literature. Instead, the previous literature has (implicitly or explicitly) assumed that f is homogenous of degree 1,  $\theta = \beta + \alpha = 1$ , and has estimated Equation [2].

In short, given two regions with an equal level of private inputs,  $(K_{i,t}, L_{i,t})$ , the region with a higher number of firms  $(n_{i,t})$  could be more productive for two reasons: first, the production has decreasing returns to scale ( $\theta < 1$ ); or second, the number of firms is a proxy or a measure of entrepreneurship capital  $E_{i,t}$ . In fact, if we assume that  $\theta = 1$  (which implies that there is no single, optimal and finite firm size), the only possible explanation is the second one. Previous literature has made the same argument and thus ruled out the returns to scale explanation.

Consequently, it could be argued that the main empirical contribution of this literature is to suggest that, in terms of regional production, it is not only the level of private inputs used at the regional level that is important, but also the number of firms among which they are distributed. Regions with a higher number of firms (smaller average firm size in terms of inputs) will be more productive.

Our proposal is to look in more depth at two possible elements affecting the total factor productivity of the firms in the region: decreasing returns to scale and regional entrepreneurship capital. A priori, we do not know which is the determining factor, and it could in fact vary between studies. In accordance with the discussion in this section, it will depend on the importance of returns to scale and the correlation between the entrepreneurship capital measure and the number of firms. In fact, this is an empirical question and we provide a methodological benchmark for addressing it: the estimation of Equation [5] is an easy first step in this direction. The following section provides some evidence in this regard.

#### 2. AN APPLICATION TO SPANISH PROVINCES

In empirical terms, the above section suggests that we should present estimations of Equation [5], as well as the estimation of Equation [1], the one estimated in the previous literature:

$$\ln \overline{Y}_{i,t} = \mu \ln R_{i,t} + \delta_5 \ln E_{i,t} + \alpha \ln \overline{K}_{i,t} + \beta \ln \overline{L}_{i,t} + \varepsilon_{i,t}$$
<sup>[5]</sup>

$$\ln Y_{i,t} = \mu \ln R_{i,t} + \delta \ln E_{i,t} + \alpha \ln K_{i,t} + \beta \ln L_{i,t} + \varepsilon_{i,t}$$
<sup>[1]</sup>

From the coefficients estimated with the two equations we can identify the effect of entrepreneurship capital on production estimated in accordance with the previous literature ( $\delta$  in Equation [1]) as the aggregation of two components: i) the presence of decreasing returns to scale ( $\delta - \delta_5$ ) and ii) the effect of entrepreneurship capital spillovers ( $\delta_5$  in Equation [5]).

For that purpose, we created panel data covering an 11-year period from 2002 to 2012 (t = 1, ..., 11), for the 52 Spanish provinces (i = 1, ..., 52), a total of 572 observations. This enables us to control for regional fixed effects and time fixed effects. The inputs and outputs included and the ways of measuring them are as similar as possible to those used in the reviewed literature. We collected this information from different sources.

The regional aggregate output is measured by the Gross Value Added  $(Y_{i,t})$ . The Spanish National Statistics Institute (INE) generates periodically disaggregated information at the provincial level on the annual value of the production of goods and services minus intermediate consumption. Like all the other monetary variables, this will be expressed in million euros (constant 2000 prices).

The BBVA Foundation and the Valencian Institute of Economic Research provide monetary values for the set of assets accumulated in each province, capital stock  $(K_{i,t})$ . Labour  $(L_{i,t})$  is measured by the number of employees engaged in production activities in each province. This is sourced from the Economically Active Population Survey periodically produced by the INE.

Following Bönte *et al.* (2008), knowledge ( $R_{i,t}$ ) is measured by the number of patents filed each year based on the available provincial-level data from the Spanish Patents and Trademarks Office. We do not have access to other proxies at the regional level used in the previous literature, such as the number of people employed in private companies or universities in areas related to R&D (Mueller 2007) and annual R&D costs (Griliches 1998). Table 1 presents descriptive statistics of the variables.

The stock of firms  $(n_{i,t})$  is required to compute firms' average production and average private inputs. This information is available from the Central Business Register (DIRCE) database. Table 2 presents descriptive statistics of this variable per region. The number of firms in one region can be considered a measure of the entrepreneurship capital of this region. Unfortunately, we do not have data for the regional start-ups for each year; therefore, we cannot provide empirical evidence using the ratio of new firms per inhabitant as in Audretsch and Keilbach (2004a,b, 2008). Entrepreneurship capital has also been measured in previous literature as the number of entrepreneurs per inhabitant (Acs *et al.* 2012). The main conclusions are similar using this ratio or the number of entrepreneurs. For simplicity's sake, we only

Table 1: Descriptive variables				
Average	Standard Deviation			
15.9535	0.9620			
17.1079	0.9935			
12.2875	0.9709			
3.1233	1.4702			
572				
	Average           15.9535           17.1079           12.2875           3.1233           572			

	Number of firms		
	Mean over years	Standard Deviation	
Alava	20,747	928	
Albacete	25,903	1,672	
Alicante	129,322	9,742	
Almeria	40,740	3,339	
Asturias	69,429	2,495	
Avila	10,894	529	
Badajoz	38,614	2,185	
Barcelona	446,137	23,567	
Burgos	24,634	1,201	
Caceres	24,915	1,681	
Cadiz	60,047	3,023	
Cantabria	37,881	1,914	
Castellon	39,740	2,703	
Ceuta	3,663	62	
Ciudad Real	30,521	1,731	
Cordova	46,716	2,518	
Coruna	79,842	3,922	
Cuenca	13,853	780	
Girona	53,007	5,190	
Granada	56,589	4,005	
Guadalajara	12,268	1,559	
Guipuzcoa	58,322	2,959	
Huelva	25,424	1,459	
Huesca	16,041	838	
Balearic Is.	87,298	4,684	

Table 2: NUMBER OF FIRMS BY PROVINCE AND FIRM SIZE

	Number of firms		
	Mean over years	Standard Deviation	
Jaen	34,947	1,641	
La Rioja	22,426	1,136	
Las Palmas	68,874	4,237	
Leon	32,382	1,167	
Lleida	34,493	2,467	
Lugo	23,982	934	
Madrid	482,953	35,165	
Malaga	106,690	8,327	
Melilla	3,683	85	
Murcia	89,650	6,925	
Navarra	41,522	1,523	
Ourense	22,830	620	
Palencia	10,706	268	
Pontevedra	66,071	3,676	
Salamanca	22,751	861	
Segovia	11,133	610	
Seville	110,561	7,978	
Soria	5,876	143	
Tarragona	53,375	3,799	
Tenerife	63,088	3,733	
Teruel	9,131	418	
Toledo	43,201	3,760	
Valencia	174,680	10,505	
Valladolid	33,771	1,904	
Vizcaya	83,581	3,714	
Zamora	12,055	321	
Zaragoza	64,566	2,376	
Total	3,181,546	183,676	
0 Employees	1,672,297	117,662	
1-9 Employees	1,335,419	80,236	
10-99 Employees	161,413	20,600	
100-499 Employees	10,697	985	
>500 Employees	1,719	121	

Table 2: NUMBER OF FIRMS BY PROVINCE AND FIRM SIZE (continuation)

present the results using the stock of firms  $(n_{i,t})$  as the measure of entrepreneurship capital<sup>1</sup>. Note that, in this case, the estimations of Equation [1] and [5] only differ in the value of the estimates of the elasticity of production with respect to entrepreneurship capital ( $\delta = 1 - \alpha - \beta + \delta_5$ ). Therefore, we are only going to present estimates from Equation [1] and infer from them  $\delta_5 = \delta + \alpha + \beta - 1$ .

Table 3 provides these estimations. In the first column, the estimated models do not include province fixed effects or time fixed effects. In the second column, only province fixed effects are included and in the last column both types of fixed effects have been included. Following the econometric literature on panel data, the group model, the province fixed effects model and the random effects model have been estimated for all the equations. The main results referring to the elasticity of production with respect to entrepreneurship capital remain unchanged. For reasons of expositional simplicity, we provide only the estimations of the fixed effects model because the Breush and Pagan (1979) and Hausman (1978) tests indicate that this is the most appropriate method for modelling the non-observable heterogeneity among provinces in the analysed sample.

Equation [1]: Dependent variable: In $Y_{i,t}$						
Independent	Coefficient	– Model 1	Model 2	Model 3		
Variables		Widdel 1		1410001 5		
Constant		4.5947***	6.0940***	13.1106***		
		[0.238]	[0.273]	[0.769]		
lnK <sub>i,t</sub>	α	0.1404***	0.2158***	0.0774**		
-,-		[0.021]	[0.016]	[0.041]		
lnL <sub>i,t</sub>	β	0.6175***	0.2103***	0.0081		
,		[0.062]	[0.033]	[0.035]		
lnR <sub>i,t</sub>	μ	0.0839***	0.0103***	0.0086**		
,		[0.009]	[0.004]	[0.003]		
lnE <sub>i,t</sub>	δ	0.1055**	0.3383***	0.1254***		
-,-		[0.060]	[0.049]	[0.048]		
Regional Effects		No	Yes	Yes		
Temporal Effects		No	No	Yes		
Observations		572	572	572		
Groups:			52	52		
R-squared		0.9841	0.7451	0.8142		

Table 3: IMPACT OF ENTREPRENEURSHIP CAPITAL ON PRODUCTIO	)N
Equation [1]: Dependent Variable: $\ln Y_{i,t}$	

\*: Significant at the 0.10 level. \*\*: Significant at the 0.05 level. \*\*\*: Significant at the 0.01 level. Clustered by province standard errors in brackets.

<sup>(1)</sup> This and all the other estimations cited in the paper but which do not appear in the text can be provided upon request to the authors.

The elasticity of production with respect to capital ( $\alpha$ ), to labour ( $\beta$ ) and to knowledge ( $\mu$ ) is positive and statistically significant at 1% in all the estimations except in model 3. When the province and time fixed effects are included, the values of these elasticities are substantially lower. In all cases, the estimated values indicate that production technology presents decreasing returns to scale<sup>2</sup> ( $\alpha + \beta < 1$ ). Previous estimations of the production function at the Spanish province level present mixed results. Pablo-Romero *et al.* (2008) report constant returns to scale for the 1990-1999 period, while Alvarez-Ayuso, *et al.* (2016) report decreasing returns to scale for the 1980-2007 period.

Elasticity of production with respect to regional entrepreneurship capital ( $\delta$ ) is in all cases positive and statistically significant at the usual levels. This result is consistent with the positive relationship found in the previous literature. Using the proposed methodology, the effect of entrepreneurship capital on production estimated in accordance with the previous literature is positive due to the presence of decreasing returns to scale ( $\theta = \alpha + \beta < 1$ ) but not to the presence of entrepreneurship capital spillovers, which according to our estimations are negative ( $\delta_5 < 0$ ). The estimated values using Equation [5] are -0.1368 in model 1 (significant at 1%), -0.2374 in model 2 (significant at 1%) and -0.7869 in model 3 (significant at 1%).

# 3. CONCLUSION AND DISCUSSION

As in the previous literature, this paper shows a positive impact of regional entrepreneurship capital on regional production. We argue that, from a theoretical point of view, the firms in those regions could be more productive for at least two reasons: because there are decreasing returns to scale or due to the fact that the number of regional entrepreneurs produces positive externalities. Previous literature on entrepreneurship capital has not properly distinguished between the two effects, so the previous evidence has only been interpreted in terms of positive externalities.

The paper presents a methodology to help to distinguish between the two effects. The approach is simple; it can be applied with data aggregated at the regional level and only requires information about the number of firms in the region. This is a starting point for analysing the sources of differences in productivity between regions as detected previously by the entrepreneurship capital literature.

Note that the proposed methodology does not centre on the measure of entrepreneurship capital used but rather the equations used to estimate its effect. The methodology suggests that the number of firms has to be used in order to control for the existence of returns to scale, but it does not claim that this is the best measure of entrepreneurship capital.

We provide evidence related with all these aspects for a sample of Spanish provinces in the 2002-2012 period. According to the estimates presented, production technologies show decreasing returns to scale in the use of private inputs: labour and capital. This seems to be the norm, and not the exception, in the reviewed literature. In this paper, this is the main explanation for the estimated positive relation-

<sup>(2)</sup> The null hypothesis that  $\beta + \alpha = 1$  is rejected at the 1% level in all equations.

ship between entrepreneurship capital and production at the regional level. According to our estimations, regional entrepreneurship capital has negative spillovers. The provided evidence cannot be taken as evidence against the knowledge spillover theory of entrepreneurship [Acs *et al.* (2009)]. This is merely a preliminary indication that the spillovers of regional entrepreneurship capital may be overestimated when we do not correct for decreasing returns to scale.

There is therefore a need for further evidence in order to confirm these results in other contexts. The literature has used other measures of entrepreneurship capital and we cannot verify exactly what would have happened in previous studies if we had applied the proposed methodology. Furthermore, we do not address other relevant issues that arise in the entrepreneurship capital literature, such as the measurement of inputs or reverse causality problems. As discussed in the theoretical sections, without information disaggregated at the firm level, it is difficult to distinguish between regional entrepreneurship capital spillovers or the existence of correlations between the size of firms and their total factor productivity. Furthermore, regional information disaggregating outputs and inputs at the sectorial level will be valuable for several reasons: i) the production functions could vary between sectors, ii) some sectors could benefit more from entrepreneurship capital spillovers and iii) sectorial entrepreneurship capital could generate different spillovers. The proposed methodology can easily be adapted to this kind of information. Indeed, it can be extended to the new theoretical or empirical relationships that have not been explored in this study. Theoretical developments may improve our understanding of the relationships between the different inputs and outputs measured. In future empirical studies, it would be useful to control for such sources of endogeneity.



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#### RESUMEN

La intervención pública para favorecer el emprendimiento se ha justificado por las externalidades que la actividad emprendedora de una región pueden tener sobre la productividad de las empresas ubicadas en dicha región. La evidencia disponible muestra una relación positiva entre la actividad emprendedora y la productividad regional. Este artículo argumenta que dicha evidencia puede ser explicada por las mencionadas externalidades, pero también por la presencia de rendimientos decrecientes a escala en las tecnologías de producción de las empresas. El artículo desarrolla un sencillo marco metodológico para medir y poder distinguir entre ambos efectos. Su aplicación a una muestra de 52 provincias españolas durante once años confirma la presencia de rendimientos decrecientes a escala. En consecuencia, la literatura previa podría estar sobre estimando las externalidades de la actividad emprendedora en la productividad de las empresas.

*Palabras clave:* capital emprendedor, productividad regional, economías de escala.

Clasificación JEL: R11, L26, O4.